使用内核方法放宽因果推断中的可观测性假设

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Talk at Causal Inference Seminars, 04.2023



Why relax observability assumptions?







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Mask interesting relationships:





Kernel Mean Embeddings (KME)



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 $\langle \mu_{P_X}, f \rangle_{H_X} = \mathbb{E}_{P_X}[f(X)]$

 $\mu_{W|a,x,z} := C_{W|A,X,Z} \left(\phi(a) \otimes \phi(x) \otimes \phi(z) \right)$



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 $\widehat{C}_{W|A,X,Z} = \operatorname*{argmin}_{C \in \mathcal{H}_{\Gamma}} \widehat{E}(C), \text{ with}$ $\widehat{E}(C) = \frac{1}{m} \sum_{i=1}^{m} \|\phi(w_i) - C\phi(a_i, x_i, z_i)\|_{\mathcal{H}_{\mathcal{W}}}^2 + \lambda \|C\|_{\mathcal{H}_{\Gamma}}^2$



$$\mu_{W|a,x,z} := C_{W|A,X,z}$$

$$\widehat{C}_{W|A,X,Z} = \underset{C \in \mathcal{H}_{\Gamma}}{\operatorname{argmin}} I$$
$$\widehat{E}(C) = \frac{1}{m} \sum_{i=1}^{m} \|\phi(w_i) - \psi(w_i)\| = \frac{1}{m} \sum_{i=1}^{m} \|\phi(w_i) - \psi(w_i)\| = \frac{1}{m} \sum_{i=1}^{m} \|\phi(w_i)\| = \frac{1}{m} \sum_{i=1}^$$

- $_{Z}(\phi(a)\otimes\phi(x)\otimes\phi(z))$
- $\widehat{E}(C)$, with
- $-C\phi(a_i, x_i, z_i)\|_{\mathcal{H}_{W}}^2 + \lambda \|C\|_{\mathcal{H}_{\Gamma}}^2$

 $\widehat{C}_{W|A,X,Z} = \Phi(W)(\mathcal{K}_{AXZ} + m \lambda)^{-1} \Phi^T(A,X,Z)$



$$\mu_{W|a,x,z} := C_{W|A,X,z}$$

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Convergence rates are well understood (Singh et al 2019, Mastouri, Zhu, et al 2021)

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- $C\phi(a_i, x_i, z_i) \|_{\mathcal{H}_{\mathcal{W}}}^2 + \lambda \|C\|_{\mathcal{H}_{\Gamma}}^2$

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Translation invariant: k(x, y) = k(x - y)



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 $\hat{\mu}[\alpha] = \hat{k}[\alpha]\psi[\alpha]$

Bochner's theorem: \hat{k} is a probability measure.

$$= k(x - y)$$

$$k(x - y)p(y)dy$$





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KRR estimate of CME: $\hat{\mu}_{X|z}^{(s)}(x) = \sum_{i=1}^{\infty} \hat{\gamma}_{j}^{(s)}(z)k(x_{j},x)$ j = 1

 $\hat{\gamma}_i^{(s)}(z) = (K_Z + s\lambda I)^{-1}K_{ZZ}$





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Fourier transform: $\tilde{\hat{\mu}}_{X|z}^{(s)}(\alpha) = \sum_{i=1}^{s} \hat{\gamma}_{j}^{(s)}(z) e^{-i\alpha x_{j}} \tilde{k}(\alpha)$





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 $(x_j, z_j)_{j=1}^s$



 $(x_j, z_j)_{j=1}^s \longrightarrow \text{Have } \hat{\mu}_{X|z}^n(y) = \sum_{j=1}^n \hat{\gamma}_j^n(z) k(x_j, y).$ i = 1

Where $\hat{\gamma}_{i}^{n}(z) = (K_{ZZ} + n\lambda^{n}I)^{-1}K_{ZZ}$.



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Where

$$\hat{\mu}^n_{X|Z} \to^n$$

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$$\hat{\mu}_{X|z}^{n}(y) = \sum_{j=1}^{n} \hat{\gamma}_{j}^{n}(z)k(x_{j}, y).$$

$$\hat{\psi}_{X|z}^n(\alpha) := \sum_{j=1}^n \hat{\gamma}_j^n(z) e^{i\alpha x_j}.$$

$$\hat{\gamma}_j^n(z) = (K_{ZZ} + n\hat{\lambda}^n I)^{-1} K_{ZZ}$$

Theorem 1. With real, translation-invariant kernel: $\mu_{X|Z}$ iff $\hat{\psi}_{X|Z}^n \to^n \psi_{X|Z}$ in IFT of kernel.

Zhu et al 2022, UAI.

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If the characteristic function of the pair (Z_1, Z_2) does not vanish, then the distribution of (Z_1, Z_2) determines the distributions of X_1 , X_2, X_3 up to a change of the location.

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$M = A + \Delta M$ $N = A + \Delta N$



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$$\underbrace{\widetilde{\mathbb{E}}_{\mathscr{P}_{A}}(\alpha) :=}_{\mathscr{P}_{A}}\left[e^{i\alpha A}\right] = \exp\left(\int_{0}^{\alpha} i \frac{\mathbb{E}\left[Me^{i\nu N}\right]}{\mathbb{E}\left[e^{i\nu N}\right]} d\nu\right)$$







Recap: Identification with instrumental variables

Identification:

 $Y = f(A) + \epsilon \quad \mathbb{E}[\epsilon | Z] = 0$ $f(A) = \mathbb{E}[Y|do(A)]$ $\mathbb{E}[Y|Z] = \int_{\mathscr{A}} f(a)p(a|Z)da$





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Identification: $Y = f(A) + \epsilon \quad \mathbb{E}[\epsilon | Z] = 0$ $f(A) = \mathbb{E}[Y|do(A)] \qquad A = \gamma Z + \epsilon_A \quad \epsilon_A \perp Z$ $\mathbb{E}[Y|Z] = \int_{\mathscr{A}} f(a)p(a|Z)da \qquad \Longrightarrow \qquad Y = \beta \gamma Z + \beta \epsilon_A + \epsilon_Y$

Linear case:

 $Y = \beta A + \epsilon_Y \quad \epsilon_Y \perp Z$





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(Strong) Assumptions: Additive error model $(Z! \perp A)_G$ $(Z \perp Y)_{G_{\bar{A}}}$

False IV: using same 'IV' for several different actions.

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???

But if $f(a) = \theta^T \phi(a)$, then simplies to $\mathbb{E}[Y|Z] = \theta^T \mathbb{E}[\phi(A)|Z]$







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???

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To induce well-posedness:

- Assume f in RKHS.
- Tikhonov regularisation.







$M = A + \Delta M$ $N = A + \Delta N$

$$\underbrace{\widetilde{\mathbb{E}}_{\mathscr{P}_{A}}(\alpha) :=}_{\mathscr{P}_{A}}\left[e^{i\alpha A}\right] = \exp\left(\int_{0}^{\alpha} i \frac{\mathbb{E}\left[Me^{i\nu N}\right]}{\mathbb{E}\left[e^{i\nu N}\right]} d\nu\right)$$





Application scenario







$$\underbrace{\psi_{\mathscr{P}_{A|z}}(\alpha) :=}_{\mathbb{P}_{A|z}\left[e^{i\alpha A} \mid z\right]} = \exp\left(\int_{0}^{\alpha} i \frac{\mathbb{E}\left[Me^{i\nu N} \mid z\right]}{\mathbb{E}\left[e^{i\nu N} \mid z\right]} dx\right)$$







How to compute the right hand side?

$$\underbrace{\psi_{\mathscr{P}_{A|z}}(\alpha) :=}_{\mathbb{P}_{A|z}\left[e^{i\alpha A} \mid z\right]} = \exp\left(\int_{0}^{\alpha} i \frac{\mathbb{E}\left[Me^{i\nu N} \mid z\right]}{\mathbb{E}\left[e^{i\nu N} \mid z\right]}dz\right)$$













Zhu et al, UAI 2022, Causal Inference with Treatment Measurement Error: A Nonparametric IV Approach.

Advantages of MEKIV

- White 2011.
- 极少的超参数调参。
- 对CME建模,而非对整个分布函数建模。



• 无关于分布的假设 对Kotlarski假设的放宽: Evdokimov and

MEKIV results



Demand Design (Mixture of Gaussians)



Summary of techniques and future work

- Kotlarski引理允许我们从它们的两个线性组合中识别三个看不 见的变量。这是否可以被继续探索?
- 特征函数和均值嵌入之间的对偶性是否可以带来更多两个方 向的融合?
- 需要放宽加性误差假设。
- 需要放宽instrumental variable假设。



Tchetgen-Tchetgen et al 2020. An Introduction to Proximal Causal Learning.



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Average causal effect estimation: $\mathbb{E}[Y|do(A = a)] = \int_{W} h(a, w, x)p(w, x)dxdw$







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Average causal effect estimation: $\mathbb{E}[Y|do(A = a)] = \int_{VW} h(a, w, x)p(w, x)dxdw$ How to get h? • Expectation operator: $\mathbb{E}[g(\cdot_U) | A, Z, X]$ • $\mathbb{E}[Y|A, U, X] = \int h(A, w, x)p(w, x | U, X)dxdw$ $\mathbb{E}[Y - h(A, W, X) | A, Z, X] = 0 \text{ a.s. } P_{AZX}$





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Average causal effect estimation: $\mathbb{E}[Y|do(A = a)] = \int_{VW} h(a, w, x)p(w, x)dxdw$ How to get h? • Expectation operator: $\mathbb{E}[g(\cdot_U) | A, Z, X]$ • $\mathbb{E}[Y|A, U, X] = \int h(A, w, x)p(w, x | U, X)dxdw$ $\mathbb{E}[Y - h(A, W, X) | A, Z, X] = 0 \text{ a.s. } P_{AZX}$ Normal regression equation: " $\mathbb{E}[Y - h(A, Z, X) | A, Z, X] = 0$ a.s. P_{AZX} " Here we also need to take the expectation over $P_{W|AZX}$.



 $\mathbb{E}[Y - h(A, X, W) | A, X, Z] = 0 \text{ a.s. } P_{AXZ}$



 $\mathbb{E}[(Y - h(A, X, W))g(A, X, Z)] = 0 \text{ a.s. } P_{AXZ} \text{ For all g}$

Mastouri*, Z.*, et al. Proximal Causal Learning with Kernels: Two-Stage Estimation and Moment Restrictions. ICML 2021.



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Precursor loss: $R(h) = \sup(\mathbb{E}[(Y - h(A, W, X))g(A, Z, X)])^2$ 8

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$\mathbb{E}[Y - h(A, X, W) | A, X, Z] = 0 \text{ a.s. } P_{AXZ}$

	•	If E[A B] = 0,
	•	Then (for g measurable):
	•	E[Ag(B)] = E[E[Ag(B) B]]
~	•	= E[E[A B]g(B)] = 0

 $\mathbb{E}[Y - h(A, X, W) | A$



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$$[X, X, Z] = 0$$
 a.s. P_{AXZ}

•	If E[A B] = 0,
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•	E[Ag(B)] = E[E[Ag(B) B]]
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$$(A, W, X))g(A, Z, X)])^2$$

• Restrict g to
$$\mathcal{H}_{\mathcal{AXE}}$$

PMMR surrogate loss $R_k(h)$ k indexes the kernel.

 $R(h) = \sup(\mathbb{E}[(Y - h(A, W, X))g(A, Z, X)])^2$ 8

 $R_k(h) =$ $g \in \mathcal{H}_{\mathscr{AZX}}, \quad \|g\| \leq 1$

 $= \mathbb{E}[(Y - h(A, W, X))(Y' - h(A', W', X'))k((A, Z, X), (A', Z', X'))]$

Precursor loss:



sup $(\mathbb{E}[(Y - h(A, W, X))\langle g, k((A, Z, X), \cdot) \rangle])^2$

Precursor loss: $R(h) = \sup(\mathbb{E}[(Y - h(A, W, X))g(A, Z, X)])^2$ *g* Restrict g to $\mathcal{H}_{\mathcal{AXF}}$ $\sup \quad (\mathbb{E}[(Y - h(A, W, X))\langle g, k((A, Z, X), \cdot) \rangle])^2$ $R_k(h) =$ $g \in \mathcal{H}_{\mathscr{AZX}}, \quad \|g\| \leq 1$ $= \mathbb{E}[(Y - h(A, W, X))(Y' - h(A', W', X'))k((A, Z, X), (A', Z', X'))]$ V-statistic: $R_V(h) := \frac{1}{n^2} \sum_{j=1}^n (y_i - h_i)(y_j - h_j)k_{ij}$ (reweighed ERM!) n^2 . i,j=1

Mastouri*, Z.*, et al. Proximal Causal Learning with Kernels: Tws Stage Estimation and Moment Restrictions. ICML 2021.